Disaster Experience Mitigates the Partisan Divide on ² Climate Change: Evidence from Texas

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Abstract

Despite the abundance of real world events and scientific information linking the wors-6 ening extreme weather to climate change, public attitudes toward climate issues in 7 the United States remain highly divided along partial lines. We compare the effect 8 of different stimuli linking extreme weather events to climate change – personal expe-9 riences and scientific information – in reducing the partian gap. A two-wave survey 10 corresponding to multiple extreme weather events in Texas, including a natural exper-11 iment with power outage data from the 2021 North American Winter Storms, shows 12 that personal experiences with extreme weather reduce the partian divide in climate 13 beliefs and polices. Scientific information attributing extreme weather events to cli-14 mate change, however, had no effect in closing the partian gap. These findings suggest 15 that extreme climate events and disaster experiences force vividly tangible information 16 about the proximity and severity of climate change on exposed individuals, prompting 17 belief-updating and preference-shifting toward pro-climate policies. 18

Keywords: climate change, environmental disasters, disaster experiences, scientific
 information, pro-environmental policy attitudes

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21 1 Introduction

Climate change-induced extreme weather events, such as wild fires in the western United 22 States and hurricanes along the Gulf Coast and Eastern Seaboard, occur with increasing 23 frequency, visibility, and consequence (Davenport, Burke and Diffenbaugh 2021; Parks and 24 Abatzoglou 2020). Experience with these extreme climate events and disasters present vividly 25 tangible stimuli about the proximity, severity, and costliness of climate change. Scientific in-26 formation attributing extreme weather and its consequences to anthropogenic climate change 27 has also become more abundant through both academic research (Trenberth, Fasullo and 28 Shepherd 2015) and public science channels (IPCC 2022). Yet, individual beliefs and policy 29 preferences about climate change in the U.S. remain deeply polarized along partial lines 30 (Leiserowitz et al. 2023; Dunlap, McCright and Yarosh 2016). This is in spite of the fact 31 that climate-skeptic individuals, who tend to be Republican, are increasingly exposed to 32 ever-growing amounts of experiential and informational stimuli about climate change. This 33 cause of partisan division is of particular importance because it is associated with gridlock 34 on climate policy-making (Hazlett and Mildenberger 2020). 35

Can extreme weather experiences and scientific information attributing extreme weather 36 to climate change reduce this partian gap? Both these *experiential stimuli* (personal expe-37 riences with extreme weather) and *informational stimuli* (scientific information attributing 38 these events to climate change) are seen to be key drivers of individuals associating climate 39 change with negative outcomes (Thaker and Cook 2021; Wong-Parodi and Garfin 2022). 40 However, despite numerous studies investigating how these two stimuli shape climate atti-41 tudes, conclusive findings about either factor have yet to be established. Empirical evidence 42 about the experiential stimuli (Howe 2021; Sisco 2021; Reser and Bradley 2020; Howe et al. 43 2019) and the informational stimuli (Rode et al. 2021) are mixed between exhibiting posi-44 tive or null effects. Moreover, scientific information even led to backfire effects among specific 45 politically-relevant subgroups (i.e., Republicans (Zhou 2016; Hart and Nisbet 2012) and cli-46 mate skeptics (Dixon, Bullock and Adams 2019; Chapman and Lickel 2016)). Recent studies 47 have begun to examine how the relationship between personal experiences and pro-climate 48 attitudes differs across political groups (Constantino et al. 2022; Hazlett and Mildenberger 49 2020; Zanocco et al. 2019; Ogunbode, Doran and Böhm 2020). Notably, Constantino et al. 50 (2022) and Zanocco et al. (2019) find evidence that negative personal experience with ex-51 treme weather decreased the partisan gap on climate attitudes, as Republicans tended to 52 shift closer to Democrats' positions. Conversely, Hazlett and Mildenberger (2020) show that 53 Republican-dominated areas in California were unresponsive to wildfire exposure when vot-54 ing on climate-policy ballots, which effectively increases the partian gap. 55

Critically, existing research does not directly compare the impacts of extreme weather 56 experiences and scientific information, two different types of stimuli prompting individuals to 57 link climate change to negative outcomes, on the same individuals. The lack of within-sample 58 comparisons leaves notable gaps in our understanding of climate attitudes. First, given sam-59 ple heterogeneity across studies, it is difficult to contextualize findings about different stimuli 60 (i.e., experiential and informational) against one another. Second, personal experiences with 61 extreme weather and scientific information on attribution is likely to conditionally impact 62 or moderate climate attitudes (Lacroix, Gifford and Rush 2020), which cannot be examined 63 unless we explicitly model the interaction effect on a sample of individuals. 64

In this paper, we fill these gaps by simultaneously examining the effects of personal expe-65 riences and scientific information in influencing the climate attitudes of partian individuals. 66 We achieve this through several research designs that we conducted as part of two-wave 67 survey (2020 and 2021) fielded in Texas, U.S., a state that has experienced both major hur-68 ricanes and extreme winter storms in recent years. Our surveys draw directly on personal 69 experiences, a preregistered experiment,¹ and a natural experiment, each measuring expo-70 sure of our survey respondents to the link between climate change and extreme weather. We 71 explored both personal experiences about hardship directly experienced from climate disas-72 ters and scientific information explicitly highlighting the link. We started with the general 73 expectation that both experiential and informational stimuli will effect pro-climate attitu-74 dinal change, then examined how the heterogeneous effects for both stimuli across partisan 75 groups can lead to a reduction in the partisan gap on a set of climate attitudes ranging from 76 belief in anthropogenic climate change to support for various pro-climate policies. 77

As previewed in the introduction of our research design above, results come from three 78 sets of analyses – survey, quasi-experimental, and experimental – that systematically explore 79 how Democrats' and Republicans' beliefs about climate change and support for pro-climate 80 policies vary by their personal experiences and exposure to scientific information. We find 81 that Republicans update their beliefs about anthropogenic climate change and climate policy 82 when they personally experience extreme weather events while Democrats generally update 83 their beliefs very little because their existing beliefs are already strongly pro-climate. The 84 observed mechanism that experiences drive pro-climate attitudes, however, also holds for 85 Democrats for outcomes not subject to a ceiling effect (i.e., their willingness to share pro-86 climate messages on social media). In terms of scientific information, experimentally provided 87 scientific attribution linking climate change and extreme weather events had no measurable 88 impact on climate change attitudes for both partian groups, even when moderated by ex-89 isting personal experiences. 90

¹See Supplementary Information S6 for our preregistration materials.

Beyond being the first study, to our knowledge, that simultaneously examines the effects 91 of different types of stimuli across a fixed set of individuals from distinct partial groups, our 92 study makes a number of additional contributions. First, we explicitly study the potential 93 for an interactive effect between the two kinds of stimuli, for which we found none. Second, 94 focusing on Texas afforded a number of benefits, most notably being able to study individ-95 uals' experiences with both expected (i.e., hurricanes) and unexpected (i.e., winter storms) 96 extreme weather events. Here, our findings are highly robust across both contexts. Third, 97 because of the timing of our surveys and the collection of real-world data, we were able 98 to measure personal experience in different ways. Specifically, we measure both perceived 99 personal experience and objective geographic exposure (i.e., being in an afflicted location at 100 the time of an extreme weather event). Perceived personal experience captures important 101 psychological realities (Reser and Bradley 2020), but it is hard to identify the causal effect of 102 perception. On the other hand, while geographic exposure - as an externally validated mea-103 sure of the state of the world – facilitates identified causal estimates, they do not perfectly 104 map onto experience as a construct (Reser and Bradley 2020) and are prone to measure-105 ment imprecision (Akerlof et al. 2013). Given the shortcomings of any singular measurement 106 approach, we opted to examine both. The results we present about the effects of personal 107 experience are weakly robust to both measurement approaches. 108

The remainder of our paper proceeds as follows. In the second section, we provide an 109 overview of our methodological approach, specifically how we measured pro-climate attitudes, 110 details about our case study, and how we implemented our surveys. In the third section, we 111 present our findings about the *experiential stimuli*, measured both as perceived personal 112 experience and as geographic exposure. In the fourth section, we present our findings about 113 the *informational stimuli*, which was embedded in a scientific information experiment. On 114 the whole, our results show that, although climate attitudes are widely viewed as inflexible, 115 especially for Republicans, individuals do update their attitudes when experiencing extreme 116 weather events. Between the two oft-examined types of stimuli prompting individuals to link 117 climate change to negative outcomes, we show that personal experiences are more effective 118 than information on scientific attribution in effecting pro-climate attitudes. 119

$_{120}$ 2 Methods

In this study, we conducted three set of analyses using data from a two-wave survey among Texas residents ($n_{wave1} = 1375$, $n_{wave2} = 305$). In this methods section, we outline methodological considerations common to all our analyses. Specifically, we discuss how we measured different facets of pro-climate attitudes, the merits and particulars of using Texas as a case study from which to draw our samples, and how we implemented our survey. Due to the variety of analyses we conducted across each of our studies, we leave the detailed discussion of each study, including how we measure different stimuli and how we made inference design choices, in each study's respective section.

¹²⁹ 2.1 Pro-climate Attitudes

To assess how widely applicable our comparison of the experiential and informational stimuli 130 is, we examined a variety of pro-climate attitudes, which are summarized in Table 1. First, 131 we included a set of general questions capturing respondents' belief in anthropogenic cli-132 mate change. Second, we asked respondents about their policy preferences, both in terms of 133 support for different approaches to climate mitigation, and in terms of support for disaster re-134 silience. Beyond these main climate attitudes, we also included two measures of social media 135 activism to capture pro-climate tendencies that have a low baseline of support across both 136 partisan groups. Finally, we included a number of additional measures that we summarize 137 in Supplementary Information S2.² 138

Concepts	Survey Measures	Wave
Belief in Anthropogenic Climate Change	Pro-climate Belief [*]	Both
Support for Climate Change Mitigation	Federal Carbon Emissions Tax	Both
	Climate Change Mitigation Spending	Both
Support for Disaster Resilience Policy	Disaster Relief Spending	Both
	Infrastructure Improvement (Flood Barrier)*	1
	Infrastructure Improvement (Power Grid)*	2
Social Media Activism	Social Media Like	1
	Social Media Retweet	1

Table 1: Measures of pro-climate attitudes.

*Additive scale measures (see Supplementary Information S5)

¹³⁹ 2.2 Texas as a case study

Texas is an ideal political and environmental context to study change to partisan beliefs about climate change. Politically, though solidly 'Red' at the state level, Texas exhibits substantial political and demographic diversity in its major metropolitan areas. Climate

²Our questionnaire is included in Supplementary Information S5.

change impacts also vary considerably by region in Texas. While Houston is at constant risk of hurricane exposure, the other metro areas are far enough from the coast that they are not directly threatened. In addition to the threat of hurricanes, Texas now faces more winter storm variation because of changes to the polar vortex. Subzero temperatures, once rare along the Gulf Coast region, are becoming more prevalent.

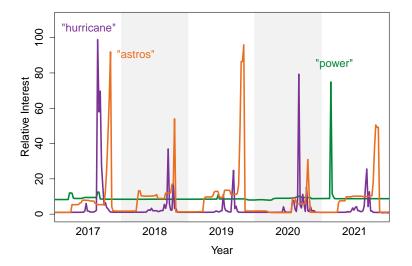


Figure 1: Comparison of relative web search interest from Texas (de-noised Google Trends) for terms associated with Hurricane Harvey, the 2021 North American winter storms, and the Houston Astros.

Further, as we show with Google Trends data in Figure 1. Texas residents have been 148 highly aware of extreme weather events and their consequences, which adds further value 149 to Texas as a case for our examination of how perceived experiences matter to pro-climate 150 attitudes. These trends explicitly capture the relative search interest on given topics within 151 Texas. Our approach is consistent with prior studies that used Google Trends to measure 152 drought awareness in California (Kam, Stowers and Kim 2019) and global interests in human 153 rights (Dancy and Fariss 2023). Major extreme whether events in Texas, such as Hurricane 154 Harvey and the 2021 winter storms, have triggered peaks in disaster awareness. Comparing 155 the relative degree of search interest for specific climate event terms to another popular 156 search term (i.e., 'astros' for Houston Astros, a highly competitive Major League baseball 157 team, which won Baseball's Major League World Series in November 2017 and played in 158 the World Series in 2019), we see the peaks of awareness in Hurricane Harvey, captured by 159 'hurricane', can be found in August–October 2017, and the peaks of awareness for the winter 160 storms, captured with searches for 'power' for power outages, are found in February 2021. 161

¹⁶² 2.3 Survey administration

We conducted a two-wave survey of Texas residents with a stated partial affiliation. The 163 first wave took place three years after Hurricane Harvey. It was was conducted between Oc-164 tober 18, 2020 and November 5, 2020, through three survey platforms, Lucid, Prolific, and 165 CloudResearch.³ Using prescreening data from each platform, we recruited Democrats and 166 Republicans who resided in Texas. We originally planned to recruit all participants using 167 Lucid, but recruitment was slow due to the constrained nature of our target population. 168 To avoid a large shift in the information environment due to election results reporting on 169 November 6, we expanded our recruitment to Prolific and CloudResearch. For these subse-170 quent samples, we implemented additional quality checks. 171

The second wave took place a few months after North American winter storms Uri and Viola in 2021. It was conducted between July 7, 2021 and October 14, 2021. For this sample, we recruited respondents from the first wave from Prolific and CloudResearch, but not Lucid because it does not support recruitment of past participants.

Field dates	Platform	n_D	\mathbf{n}_R	Remuneration
Wave 1				
Oct. 18 – Oct. 23, 2020	Prolific	96	72	\$2
Oct. 24 – Nov. 5, 2020	Lucid	424	380	up to $$4$
Oct. 29 – Nov. 5, 2020	Prolific	172	81	\$2
Oct. 30 – Nov. 5, 2020	CloudResearch	87	63	\$2
Wave 2				
Jul. 7 – Aug. 30, 2021	Prolific	116	62	\$2
Aug. 31 – Oct. 14, 2021	Prolific	42	25	\$4
Sep. 24 – Oct. 14, 2021	CloudResearch	36	24	\$2

Table 2: Survey recruitment details by wave.

 \mathbf{n}_D and \mathbf{n}_R respectively indicate sample size of Democrats and Republicans.

In both Waves 1 and 2, at the beginning of the study, participants were given a consent form that described the study instrument (i.e., answering questions on demographics and disaster experiences, reading a news article about disasters), ensured that their responses will be kept anonymous, and that the study involved minimal risks. After the study, participants were debriefed with the purpose of the study (i.e., better understand how citizens are affected by disasters and evaluate political issues), and were provided with the contact information of the study team. The Wave 1 survey took approximately 12 minutes to complete and the

³Prior to the launch, we conducted a pilot on Lucid with 132 respondents (74 Democrats and 59 Republicans) who are not included in the final data set due to mismatches with our sampling criteria and other data quality concerns (i.e., speeders or spammers). Based on the pilot, we implemented more quality controls for the full launch.

¹⁸³ Wave 2 survey took approximately 8 minutes to complete.

In the first wave, a total of 1375 eligible respondents (779 Democrats and 596 Republicans) were included in the analysis. In the second wave, the sample consisted of 305 respondents (194 Democrats and 111 Republicans) who participated in the first wave. The 305 Wave 2 respondents equate to a 53.4% retention of the subset of Wave 1 respondents we recruited for our Wave 2 survey. A full breakdown of the participant pool by survey platform and partisanship is in Table 2.⁴

¹⁹⁰ 2.4 Data availability, analysis, and results reproduction

All analysis for our study was conducted in R v4.2.2 (R Core Team 2022). Estimation for the difference-in-differences models was done with the fixest v0.11.1 package (Bergé 2018). All marginal effect calculations were done with the marginaleffects v0.9.0 package (Arel-Bundock 2023). All reproduction code will be made publicly available under the MIT license at [Github repository redacted]. All study data will be made publicly available under the CC BY 4.0 license at [Zenodo repository redacted].

¹⁹⁷ **3** Personal Experience with Extreme Weather Events

Personal experiences is difficult to measure, and any singular measurement approach has its shortcomings. We therefore opted to examine personal experience in drastically different but complementary ways: as perceived personal experience and as externally-validated geographic exposure.

202 3.1 Perceived Personal Experiences with Extreme Weather

To measure perceived personal experience with Hurricane Harvey, which caused severe dam-203 age in southeast Texas in August 2017, we asked participants in the first wave of our survey 204 whether they were personally harmed by Hurricane Harvey on three dimensions, personal 205 health, financial situation, and property damage. In the second wave, we similarly measured 206 perceived personal experience with the 2021 winter storms with a set of fourteen ques-207 tions about whether they experienced different negative events during the winter storms, 208 including perceived danger, injury, and property damage (adapted from Harville, Jacobs 200 and Boynton-Jarrett 2015). For both waves, we summed responses from the different ques-210 tions then rescaled them to the unit interval to obtain our measure of perceived personal 211

 $^{^{4}}$ A breakdown of the distribution of basic sociodemographic variables for our Wave 1 and Wave 2 surveys is in Supplementary Information S1.

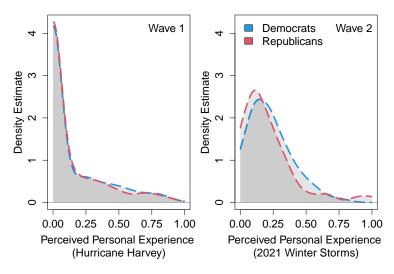


Figure 2: Distribution of perceived personal experience with Hurricane Harvey (Wave 1) and the 2021 North American winter storms (Wave 2) in Texas, rescaled to the unit interval.

 212 experience.⁵

Figure 2 shows the distribution of the self-reported exposure for both waves by partisanship, which illustrate that while our results differed by respondent's partisanship, it is not due to differences in their perceived experiences.

To test whether perceived personal experiences with extreme weather promote proclimate attitudes, we fit linear models that examine how various climate attitudes are associated with our measure. Further, to examine how partisan identity moderates the relationship between perceived personal experience and climate attitudes, we included an interaction term between partisanship and experience in the models. We also included a set of individual-level control variables in all models: ideology, age, gender, education, and indicators for Hispanic and Black identification.

We find a large difference between Republicans and Democrats (Figure 3). In general, among Republicans, perceived personal experience with both Hurricane Harvey (Wave 1) and the 2021 winter storms (Wave 2) are positively and statistically significantly associated with pro-climate attitudes. Specifically, with the single exception of beliefs about anthropogenic climate change in Wave 1, responses indicating more experience with disaster damages is predictive of greater support for both climate change mitigation and disaster resilience policies.⁶

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In contrast, among Democrats, there is no statistically discernible relationship between

⁵Additional information on our perceived personal experience measures, including distributional breakdowns, are in Supplementary Information S1.

 $^{^{6}}$ We show in Supplementary Information S3 that subsetting the Wave 1 analysis to only respondents retained in Wave 2 yields similar results. We also discuss evidence that alleviates concerns about selection bias for Wave 2 results.

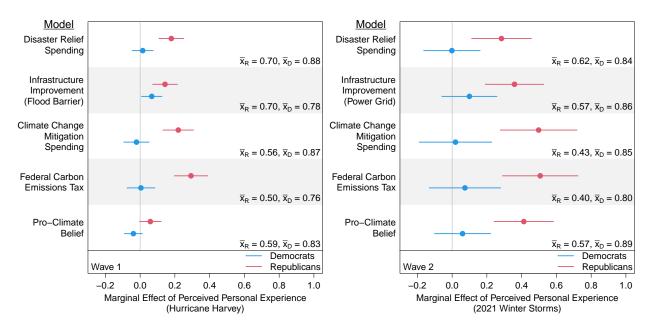


Figure 3: Relationships between perceived personal experience and climate attitudes (point estimates and 95% CIs), for Wave 1 survey respondents (left) and for Wave 2 survey respondents (right). \bar{x}_R and \bar{x}_D refer to, respectively, the sample mean of the outcome variable for the Republican and Democrat groups.

perceived personal experience and our outcomes. While this discrepancy may appear coun-231 terintuitive, additional tests show that the null finding among Democrats can be attributed 232 to a ceiling effect (Gillis et al. 2023; Zanocco et al. 2019), whereby many Democrats already 233 possess high levels of pro-climate beliefs – see Democrat group means \bar{x}_D in Figure 3 – and 234 therefore cannot increase their support. In anticipation of this potential ceiling effect, we 235 included in Wave 1 two items on willingness to share pro-climate information on social me-236 dia, which tends to have a low baseline tendency among both partisan groups. We asked 237 respondents how likely they are to retweet and to 'like' on Twitter a pro-climate mitigation 238 report.⁷ both of which are costly public acts of engagement. 239

As expected, as shown in Figure 4, because the baseline tendency to engage in social media activism is generally low, we do not observe the ceiling effect for Democrats. Instead, we find a positive relationship between perceived personal experiences and social media activism for both partisan groups. This finding suggests that the mechanism underlying the relationship between personal experience and pro-climate attitudes is similar across partisan lines.

⁷Self-reported willingness to share information on social media tends to predict observed retweeting patterns (Mosleh, Pennycook and Rand 2020).

²⁴⁶ 3.2 Natural Experiment of Geographic Exposure to the 2021 ²⁴⁷ Winter Storms

In February 2021, three months after we fielded our first survey, two overlapping winter storms (Uri and Viola) struck various parts of North America, including Texas. The timing of this event, occurring right before our Wave 2 survey, allows us to implement a convincing pretest-posttest design to examine the causal effect of geographic exposure to the winter storms as the treatment in a natural experiment.

For this study, we measured geographic exposure to the winter storms, which is an ex-253 ternally validated measure of exposure, as the extent to which individuals experienced power 254 outages during mid-late February 2021. We estimated this using data from PowerOutage.US, 255 a data aggregation company that tracks outage reports from utility companies in the U.S. In 256 Texas, this comprised raw data from 62 utility providers tracking the accounts of 13.4 million 257 customers. We aggregated the raw data (counts of outages and non-outages by geographi-258 cal area) to the city level or county level depending on data availability as the proportion 259 of customers exposed to outage during the specified time period. Specifically, counties ex-260 ceeding a certain proportion of tracked-but-not-geolocated households are aggregated to the 261 county level whereas counties with city-level data exceeding the information threshold were 262 kept at the more precise city level. We refer to this hybrid-level geographical unit as the 263 ZIP-associated region.⁸ Then, using respondents' self-reported ZIP codes, we matched them 264 to the average power outage in their ZIP-associated region during the February 13–21 period 265 which we take as our measure of geographic exposure treatment. Figure 5 shows that Texas 266 residents experienced unusually high levels of outages when the storms hit in February 2021 267 compared to February 2020. 268

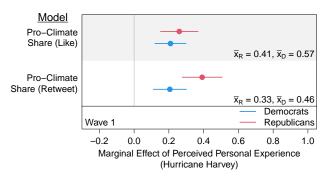


Figure 4: Relationships between perceived personal experience and willingness to share pro-climate information on social media (point estimates and 95% CIs), for Wave 1 survey respondents. \bar{x}_R and \bar{x}_D refer to, respectively, the sample mean of the outcome variable for the Republican and Democrat groups.

⁸See Supplementary Information S3 for evidence that our main findings (Figure 6), which was based on a 25% threshold, are robust to thresholds ranging from 5–45%.

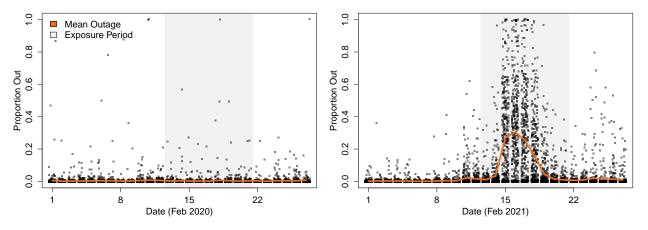


Figure 5: Proportion of households experiencing power outage by tracked administrative unit (i.e., counties or cities) in Texas during February 2020 (left) and during the winter storm in February 2021 (right).

Using this geographic exposure treatment variable and outcomes from our surveys, we used a generalized difference-in-differences design to estimate the impact of geographic exposure to extreme weather events on individuals' climate beliefs and policy preferences. As before, we consider how this effect varies by partial policy preferences term between the treatment variable and partial policy. We fit the following linear regression model:

$$Y_{izt} = \alpha_i + \tau_t + \gamma(outage_z \times storm_t) + \delta(democrat_i \times outage_z \times storm_t) + \epsilon_{izt}, \quad (1)$$

where Y_{izt} is the belief or attitude of individual *i* at time *t*, and *z* indicates the ZIP-associated 269 region individuals reside in. $outage_z \times storm_t$ is the treatment of the 2021 winter storms. We 270 are interested in the difference between Republicans and Democrats, so we further interacted 271 the treatment with partial partial (i.e., the *democrat* indicator). γ and $\gamma' \equiv \gamma + \delta$ therefore 272 capture, respectively, the treatment effects for Republicans and Democrats. We additionally 273 included in our model individual and time fixed effects (α_i and τ_t). Because the treatment 274 was assigned to the geographical unit, we conducted the analysis using standard errors that 275 were clustered at the level of the administrative unit. 276

Figure 6 shows the treatment effects of geographic exposure to power outage during the 9-day period when Texas was hit by the winter storms (February 13–21, 2021). We find that, on balance, the effect of geographic exposure to power outages on climate attitudes is much weaker than the effect we found for perceived personal experience to the winter storms. Among Republicans, for whom perceived personal experience strongly predicts greater support for all tested climate mitigation and disaster resilience policies, geographic exposure to power outages only affects preferences toward disaster relief spending.

Additional evidence (see Supplementary Information S4) suggests that our null findings are attributable to the low precision in the operationalized measure of exposure to power out-

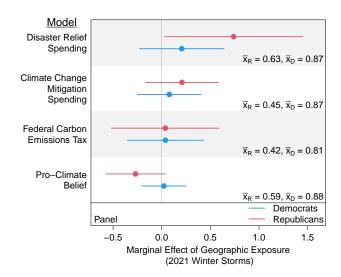


Figure 6: Treatment effects of geographic exposure to the 2021 power outages on climate attitudes (point estimates and 95% CIs), using a panel design for survey respondents who participated in both Wave 1 and Wave 2 surveys.

age – in line with prior work showing that individuals only accurately perceive very localized
extreme weather (Akerlof et al. 2013) – and would otherwise be stronger if exposure could
be measured with greater precision at the individual level. Specifically, our ZIP-associated
regions are large and there is likely to be non-negligible variation in power outages within a
region, presenting a type of measurement error that should bias the estimated effect toward
zero.

²⁹² 4 Scientific Information Experiment with Attribution of ²⁹³ Winter Storms to Climate Change

To examine whether scientific information that attribute extreme weather and its costs to 294 climate change reduces the partian divide on climate attitudes, we embedded an experi-295 ment in Wave 2 of our survey that emphasizes the link between the winter storms' extreme 296 southward extension and climate change.⁹ Specifically, Wave 2 respondents were randomly 297 assigned with equal probability into treatment and control conditions, where the former 298 were exposed to the highlighted portions of Figure 7 that explain the link between raising 299 temperatures in the arctic and extreme winter storms in Texas. To standardize respondent 300 familiarity with the winter storms, the baseline (unhighlighted) portions outlining the out-301 come of recent extreme weather events in Texas were shown in both conditions. 302

To account for the possibility of failure in experimental stimuli uptake due to respondent inattention, we implemented a number of treatment validation checks. First, we included a

⁹Supplementary Information S6 contains our preregistration plan.

Dangers of natural disasters in Texas<mark>: The role of climate change</mark>

Hurricanes have exposed Texas to the threat of disaster every year. In recent years, Texas has been affected by major hurricanes, such as Rita in 2015, Harvey in 2017, and Laura in 2020, causing countless deaths and billions of dollars in property damage annually.

The recent winter storm posed another kind of natural disaster threat to Texas. At least 57 people died in Texas as a result of the recent winter storm, according to the state health agency. The winter storm caused Texas to experience subfreezing temperatures and overwhelmed the state's electricity infrastructure, causing massive power outages. At the height of the crisis, nearly 4.5 million Texas homes and businesses were without power.

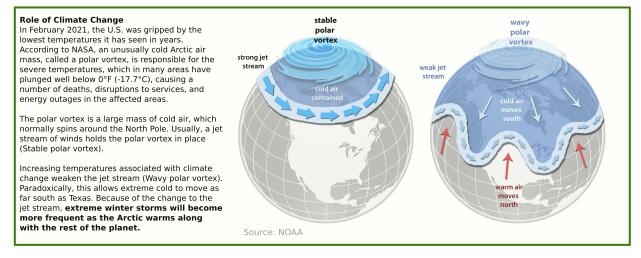


Figure 7: Experimental stimuli from the scientific information study. Parts highlighted in green are shown to the treatment group only, while unhighlighted parts are shown to treatment and control groups. (Diagram obtained from the National Oceanic and Atmospheric Administration (2019).)

manipulation check question after presenting the respondents with the scientific information. 305 This question asked respondents which of the following describes what the report they read 306 was about: 1) recent natural disasters, 2) recent natural disasters and scientific explanation 307 for winter storms, 3) recent natural disasters and the COVID-19 pandemic, or 4) the COVID-308 19 pandemic; respondents could also answer that they 5) do not know. Across the two 309 conditions, 90% of the control group correctly answered with response 1 (10% chose response 310 2), and 93% of the treatment group correctly answered with response 2 (7% chose response 311 1). These results indicate a very high rate of compliance with our treatment in terms of 312 understanding the scientific information presented. 313

Next, we checked how long the respondents spent reading the scientific information, measured in terms of how long they were on the questionnaire page containing the experimental stimuli. In median times, respondents from the control group spent approximately 33 seconds (23 and 44 seconds for the first and third quartiles), and those from the treatment group, who were shown a much longer experimental stimuli, spent approximately 72 seconds (45 and 122 seconds for the first and third quartiles). On the whole, the time our respondents spent on the stimuli page is in line with our expectations for how long they should spend.

We proceed with our analysis as it appears that the respondents took reasonable care in processing the experimental stimuli. We fit linear models where the effect of the treatment variable (i.e., scientific attribution of extreme weather to climate change) on support for

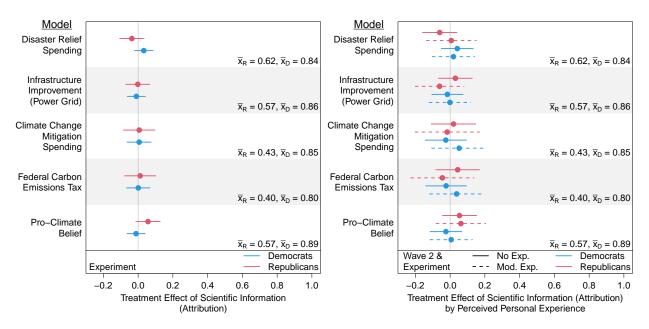


Figure 8: Treatment effect of scientific information attributing extreme weather to climate change (point estimates and 95% CIs), for Wave 2 survey respondents (left), and the same effects moderated by respondents' perceived personal experiences (right). \bar{x}_R and \bar{x}_D refer to, respectively, the sample mean of the outcome variable for the Republican and Democrat groups.

pro-climate attitudes varies by respondent partisanship. Figure 8 shows that the scientific information treatment has no discernible effect on pro-climate attitudes. Across all models, the difference between the treatment and control conditions is indistinguishable from zero, both in terms of statistical significance and substantive effect.

Finally, to test whether uptake of scientific information depends on existing personal 328 experiences, we fit additional models that let the treatment effect of scientific information 329 vary with the respondent's perceived personal experience with the 2021 winter storms. As 330 we show in Figure 8, the scientific information treatment still has no effect when subsetting 331 by respondents' personal experiences. Based on likelihood ratio tests, the expanded model 332 (i.e., interaction between scientific treatment and perceived personal experience) and reduced 333 model (i.e., without interaction term) are statistically indistinguishable from each other for 334 all outcome variables. 335

336 5 Discussion

There is an ever-growing amount of experiential stimuli and informational stimuli that prompts individuals to link the costs of extreme weather to climate change. Leveraging Texas's exposure to Hurricane Harvey in 2017 and the North American winter storms in 2021, we used a two-wave survey of Texas residents to simultaneously examine and compare the effect of personal experiences with extreme weather events and the effect of scientific information attributing these events to climate change. Across a set of survey, quasi-experimental,
and experimental results, we show that personal experiences shape people's belief in anthropogenic climate change and support for pro-climate policies but scientific information does
not.

Measuring the first stimuli, personal experience with extreme weather, in two ways, we 346 find that self-reported perceived personal experience was substantially and consistently as-347 sociated with pro-climate attitudes in various forms while externally-validated geographic 348 exposure to power outages during the 2021 winter storm exhibited weaker, but causally-349 identified, effects. Due to what are likely ceiling effects for Democrats, the effect of personal 350 experiences differed by partian groups, which led to an overall closing of the partian gap. 351 With our outcome and independent variables rescaled to the unit interval, the effect of per-352 ceived personal experience for Republicans, averaged across all main outcomes, is 0.16 for 353 Hurricane Harvey and 0.41 for the 2021 winter storms, and statistically significant for all out-354 comes but one. These effects are, respectively, approximately 33% and 105% increases from 355 the baseline averages of when Republicans have no perceived extreme weather experience. 356 In real-world terms, this means that if everyone in our sample were to perceive the highest 357 level of personal experience with Hurricane Harvey and with the winter storms, Republicans 358 pro-climate beliefs would be at approximately 92% and 96% of Democrats' levels. 359

On the other hand, the effect of scientific information is indistinguishable from zero, 360 both in terms of statistical significance and substantive effect, for all outcomes regardless of 361 whether we included existing personal experiences as a moderator. Given the evidence we 362 provided about respondents reasonably engaging with our treatment stimuli, why is there 363 no effect? Prior work on partial and climate beliefs have identified different types of 364 goals in information processing that could lead to Republican individuals rejecting scientific 365 information (Bayes and Druckman 2021; Druckman and McGrath 2019). Directional goals 366 (commonly referred to as motivated reasoning) induce individuals to resist belief updating 367 that runs counter to their priors, whereas accuracy goals, when coupled with distrust of the 368 outgroup (e.g., liberal scientific elites (Sarathchandra and Haltinner 2021)), would result in 369 stronger belief in climate skeptic information from ingroup (i.e., Republican) elites. 370

In our additional analyses, reported in Supplementary Information S2, we examined how trust in climate science varied by personal experience with extreme weather. This outcome behaves similar to others we examined, with Republicans who report higher personal experiences exhibiting greater trust toward climate science and scientists.¹⁰ This finding has implications for understanding information processing in response to climate impacts, but

¹⁰The trust in science questions were asked before respondents were assigned to the scientific information experimental conditions, so there is no treatment effect.

additional research is needed. Given that we generally still found scientific information to be ineffective despite Republicans showing attitude change toward trust in climate science, it could mean that directional motivated reasoning strongly dominates information processing for Republicans. Alternatively, we might interpret the relationship between personal experiences and increased trust in science as evidence of strong accuracy goals in information processing, where the null effect of scientific information is due to the weakness of our stimulus.

This suggests a fruitful direction for future research seeking to adjudicate between dif-383 ferent types of informational processing goals of simultaneously examining trust toward 384 different actors (e.g., scientists and ingroup elites) while varying the strength of the scientific 385 information stimulus. In our experiment, we attempted to strengthen the scientific informa-386 tion stimulus using graphical cues. Could we perhaps strengthen this visual cue with, for 387 example, a short video? Would a full semester undergraduate course be required instead? 388 We cannot answer these questions with our current study design but examining whether 389 stronger treatments could successfully convey scientific information is a promising avenue 390 for future research. 391

392 6 Conclusion

What then do our results suggest for other settings and samples? Because extreme weather 393 events are increasingly visible and experienced across the U.S., our expectation is that a 394 growing number of individuals across all states are perceiving these experiences. While we 395 expect that a ceiling effect also exists for Democrats in the broader population because they 396 are likely to have already high support for pro-climate policies, Republicans can still shift 397 their policy preferences toward greater climate policy support. However, consistent with 398 our results on social media behaviors, we expect both Democrats and Republicans to be 399 still capable of changing beliefs and behaviors as they experience climate change-related 400 events. Generally, we speculate that our findings will hold for similarly situated Democrats 401 and Republicans in other states even though we are cognizant of the fact that our Texas 402 sample is not necessarily representative of the U.S. population. We nonetheless have a clear 403 predication for individuals in other states, and we believe extensions of our results to other 404 contexts to be an important area for future research. 405

Another question related to the generalizability of our findings going forward pertains to the accumulation of experiences with climate change-related events. In July 2024, Hurricane Beryl swept across the Caribbean, Mexico, Texas, and much of the Midwestern and Northeastern U.S., setting the record as the earliest Category 5 storm observed in the Atlantic science-based informational stimuli more broadly.

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in a given year. Even as its strength waned over land, Hurricane Beryl brought flash floods 410 and tornadoes. This is just the latest in a continuing string of increasingly frequent and 411 severe climate events. How do cumulative experiences with these events impact individuals? 412 behaviors and policy preferences? This is an important and open question for future research. 413 Relatedly, personal experience with the Texas winter storm in 2021 differs from regular 414 experiences with hurricanes along with Gulf Coast of Texas in that it is a relatively novel 415 phenomenon. What does this mean for policy beliefs and behavioral changes? Though we 416 find no evidence that scientific information changes beliefs, its effects might actually vary by 417 the mix of how novel and how unfamiliar the information is perceived to be. Further work 418 should look to systematically compare different types of scientific attribution, and even other

All of these suggestions for future research would benefit from a multi-wave, panel design 421 that draws a representative sample from across the U.S., which could build on our design 422 and some of the suggestions we have made here. This would be an expensive undertaking 423 but an essential one for both the scientific community and the policy community. In terms 424 of policy implications right now, the results are clear: individuals are supportive of policies 425 that address the effects of climate change when they have experienced climate change related 426 events. Because this experience closes the partian gap, policy makers should be able to 427 generate bipartisan support on policy solutions on extreme weather events even when there 428 is ideological disagreement about climate change itself. 429

Overall, our study adds to the nascent body of research indicating that under the right 430 conditions, personal experience with extreme weather or disasters can bridge the partian 431 gap on climate attitudes (Constantino et al. 2022; Zanocco et al. 2019). We identified a 432 context in which Republicans update their beliefs about climate change and climate policy 433 preferences in response to personally-experienced climate threats. However, questions remain 434 as to whether these effects are strong enough to translate to policy-relevant behavior such 435 as voting, and whether the relative strength between experiential and informational stimuli 436 will hold under different contexts. 437

438 7 References

- Akerlof, Karen, Edward W Maibach, Dennis Fitzgerald, Andrew Y Cedeno and Amanda Neuman. 2013. "Do
 people "personally experience" global warming, and if so how, and does it matter?" *Global environmental change* 23(1):81–91.
- Arel-Bundock, Vincent. 2023. marginal effects: Predictions, Comparisons, Slopes, Marginal Means, and Hy pothesis Tests. R package version 0.9.0.
- Bayes, Robin and James N Druckman. 2021. "Motivated reasoning and climate change." Current Opinion
 in Behavioral Sciences 42:27–35.
- Bergé, Laurent. 2018. "Efficient estimation of maximum likelihood models with multiple fixed-effects: the R
 package FENmlm." CREA Discussion Papers 13.
- Chapman, Daniel A and Brian Lickel. 2016. "Climate change and disasters: How framing affects justifications
 for giving or withholding aid to disaster victims." Social psychological and personality science 7(1):13–20.
- Constantino, Sara M, Alicia D Cooperman, Robert O Keohane and Elke U Weber. 2022. "Personal hard ship narrows the partisan gap in COVID-19 and climate change responses." *Proceedings of the National Academy of Sciences* 119(46):e2120653119.
- ⁴⁵³ Dancy, Geoff and Christopher J. Fariss. 2023. "The Search for Human Rights: A Global Analysis Using
 ⁴⁵⁴ Google Data." American Political Science Review 118(1):252–273.
- ⁴⁵⁵ Davenport, Frances V, Marshall Burke and Noah S Diffenbaugh. 2021. "Contribution of historical precipita ⁴⁵⁶ tion change to US flood damages." *Proceedings of the National Academy of Sciences* 118(4):e2017524118.
- ⁴⁵⁷ Dixon, Graham, Olivia Bullock and Dinah Adams. 2019. "Unintended Effects of Emphasizing the Role of
 ⁴⁵⁸ Climate Change in Recent Natural Disasters." *Environmental Communication* 13(2):135–143.
- ⁴⁵⁹ Druckman, James N. and Mary C. McGrath. 2019. "The evidence for motivated reasoning in climate change
 ⁴⁶⁰ preference formation." *Nature Climate Change* 9(2):111–119.
- ⁴⁶¹ Dunlap, Riley E, Aaron M McCright and Jerrod H Yarosh. 2016. "The political divide on climate change:
 ⁴⁶² Partisan polarization widens in the US." *Environment: Science and Policy for Sustainable Development*⁴⁶³ 58(5):4-23.
- Gillis, Ash, Nathaniel Geiger, Kaitlin Raimi, Julia Lee Cunningham and Melanie A Sarge. 2023. "Climate
 change-induced immigration to the united states has mixed influences on public support for climate change
 and migrants." Climatic Change 176(5):48.
- Hart, P Sol and Erik C Nisbet. 2012. "Boomerang effects in science communication: How motivated reasoning
 and identity cues amplify opinion polarization about climate mitigation policies." *Communication research*39(6):701–723.
- Harville, Emily W., Marni Jacobs and Renée Boynton-Jarrett. 2015. "When Is Exposure to a Natural
 Disaster Traumatic? Comparison of a Trauma Questionnaire and Disaster Exposure Inventory." *PLOS*ONE 10(4):e0123632.
- Hazlett, Chad and Matto Mildenberger. 2020. "Wildfire exposure increases pro-environment voting within
 democratic but not republican areas." *American Political Science Review* 114(4):1359–1365.
- Howe, Peter D. 2021. "Extreme weather experience and climate change opinion." Current Opinion in Be havioral Sciences 42:127–131.

- Howe, Peter D, Jennifer R Marlon, Matto Mildenberger and Brittany S Shield. 2019. "How will climate change shape climate opinion?" *Environmental Research Letters* 14(11):113001.
- ⁴⁷⁹ IPCC. 2022. Climate Change 2022: Impacts, Adaptation and Vulnerability. Contribution of Working Group
 ⁴⁸⁰ II to the Sixth Assessment Report of the Intergovernmental Panel on Climate Change. Cambridge, UK
- 481 and New York, NY, USA: Cambridge University Press.
- 482 URL: https://report.ipcc.ch/ar6/wg2/IPCC_AR6_WGII_FullReport.pdf
- Kam, Jonghun, Kimberly Stowers and Sungyoon Kim. 2019. "Monitoring of Drought Awareness from Google
 Trends: A Case Study of the 2011–17 California Drought." Weather, Climate, and Society 11(2):419–429.
- Lacroix, Karine, Robert Gifford and Jonathan Rush. 2020. "Climate change beliefs shape the interpretation
 of forest fire events." *Climatic Change* 159:103–120.
- Leiserowitz, A., E. Maibach, S. Rosenthal, J. Kotcher, E. Goddard, M. Ballew, J. Marlon, M. Verner, S. Lee,
 J. Carman, T. Myers, M. Goldberg and N. Badullovich. 2023. *Climate Change in the American Mind: Politics & Policy, Spring 2023.* Yale Program on Climate Change Communication.
- Mosleh, Mohsen, Gordon Pennycook and David G Rand. 2020. "Self-reported willingness to share political
 news articles in online surveys correlates with actual sharing on Twitter." *Plos one* 15(2):e0228882.
- National Oceanic and Atmospheric Administration. 2019. The Science Behind the Polar Vortex.
 URL: https://twitter.com/NWSWPC/status/1090287763512049665
- Ogunbode, Charles A, Rouven Doran and Gisela Böhm. 2020. "Individual and local flooding experiences
 are differentially associated with subjective attribution and climate change concern." *Climatic Change* 162:2243–2255.
- Parks, Sean A and John T Abatzoglou. 2020. "Warmer and drier fire seasons contribute to increases in area burned at high severity in western US forests from 1985 to 2017." *Geophysical Research Letters* 47(22):e2020GL089858.
- ⁵⁰⁰ R Core Team. 2022. R: A Language and Environment for Statistical Computing. Vienna, Austria: R Foun ⁵⁰¹ dation for Statistical Computing.
- Reser, Joseph P and Graham L Bradley. 2020. "The nature, significance, and influence of perceived personal
 experience of climate change." Wiley Interdisciplinary Reviews: Climate Change 11(5):e668.
- Rode, Jacob B, Amy L Dent, Caitlin N Benedict, Daniel B Brosnahan, Ramona L Martinez and Peter H
 Ditto. 2021. "Influencing climate change attitudes in the United States: A systematic review and meta analysis." Journal of Environmental Psychology 76:101623.
- Sarathchandra, Dilshani and Kristin Haltinner. 2021. "A Survey Instrument to Measure Skeptics' (Dis)Trust
 in Climate Science." *Climate* 9(2):18.
- Sisco, Matthew Ryan. 2021. "The effects of weather experiences on climate change attitudes and behaviors."
 Current Opinion in Environmental Sustainability 52:111–117.
- Thaker, Jagadish and Christopher Cook. 2021. "Experience or attribution? Exploring the relationship be tween personal experience, political affiliation, and subjective attributions with mitigation behavioural
 intentions and COVID-19 recovery policy support." Journal of environmental psychology 77:101685.
- Trenberth, Kevin E, John T Fasullo and Theodore G Shepherd. 2015. "Attribution of climate extreme events." *Nature Climate Change* 5(8):725.
- Wong-Parodi, Gabrielle and Dana Rose Garfin. 2022. "Hurricane adaptation behaviors in Texas and Florida:
 exploring the roles of negative personal experience and subjective attribution to climate change." Envi-
- ⁵¹⁸ ronmental Research Letters 17(3):034033.

519 Zanocco, Chad, Hilary Boudet, Roberta Nilson and June Flora. 2019. "Personal harm and support for

- climate change mitigation policies: Evidence from 10 US communities impacted by extreme weather."
 Global Environmental Change 59:101984.
- Zhou, Jack. 2016. "Boomerangs versus javelins: how polarization constrains communication on climate
 Environmental Politics 25(5):788–811.